

Using Moving Horizon Estimation to Overcome Extended Kalman Filtering Failure

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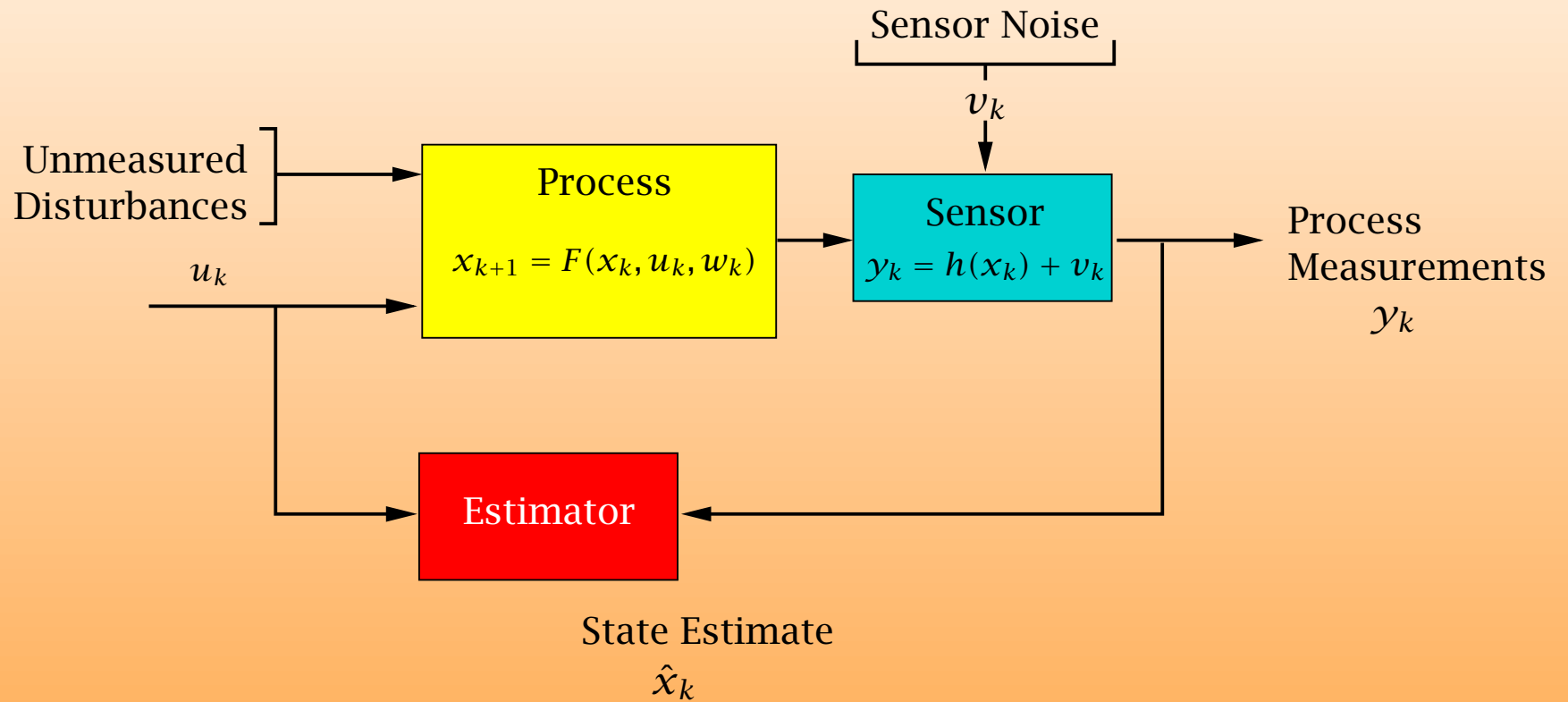
TWMCC

19 November 2003

Outline

- State estimation overview
 - problem formulation
 - extended Kalman filter
 - moving horizon estimation
- Effect of arrival cost
- Closed-loop control and plant-model mismatch
- Conclusions

State Estimation Overview



What estimate is desired?

- We will consider the model

$$\mathbf{x}_{k+1} = F(\mathbf{x}_k, \mathbf{u}_k) + G\mathbf{w}_k$$

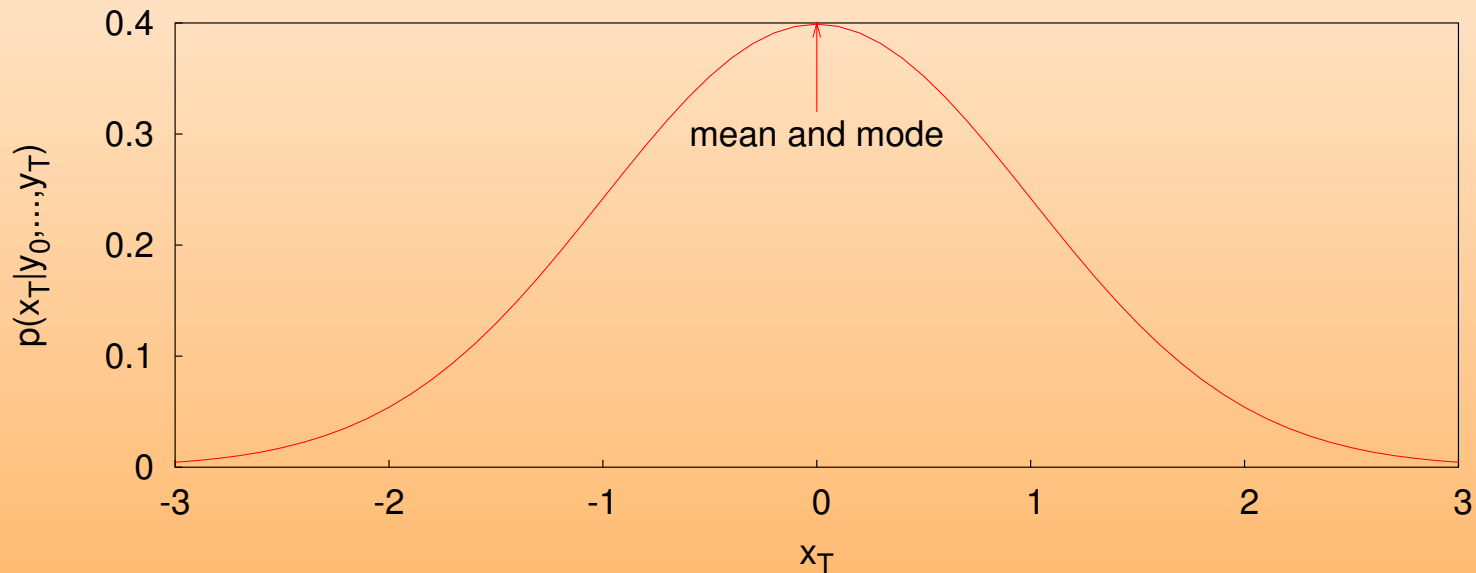
$$\mathbf{y}_k = h(\mathbf{x}_k) + \mathbf{v}_k$$

- If we knew the a posteriori distribution

$$p(\mathbf{x}_T | \mathbf{y}_0, \dots, \mathbf{y}_T)$$

what point estimate should we calculate?

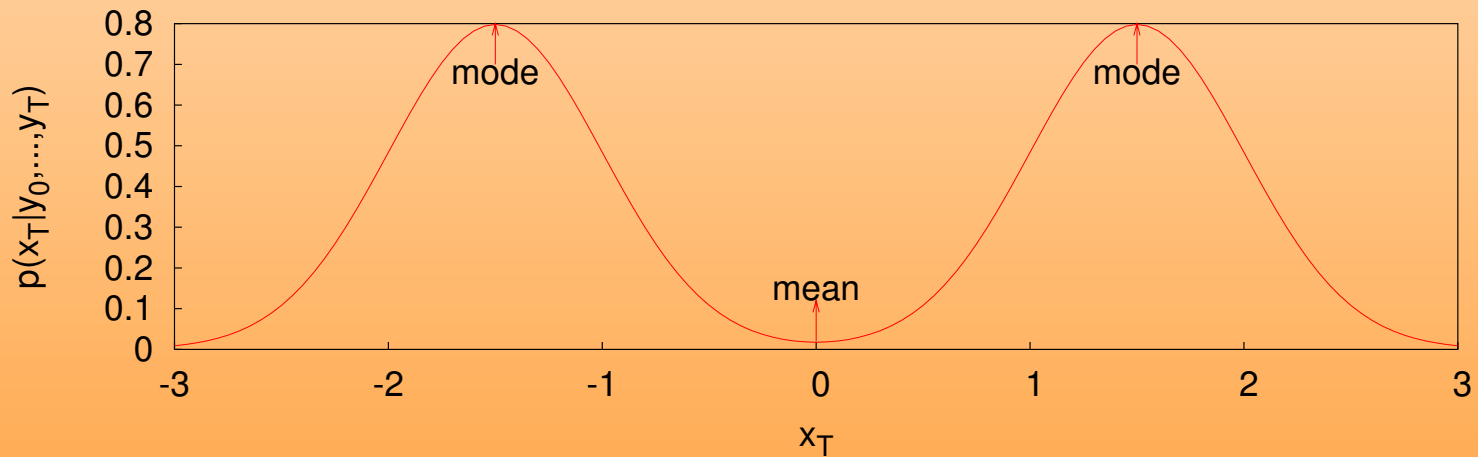
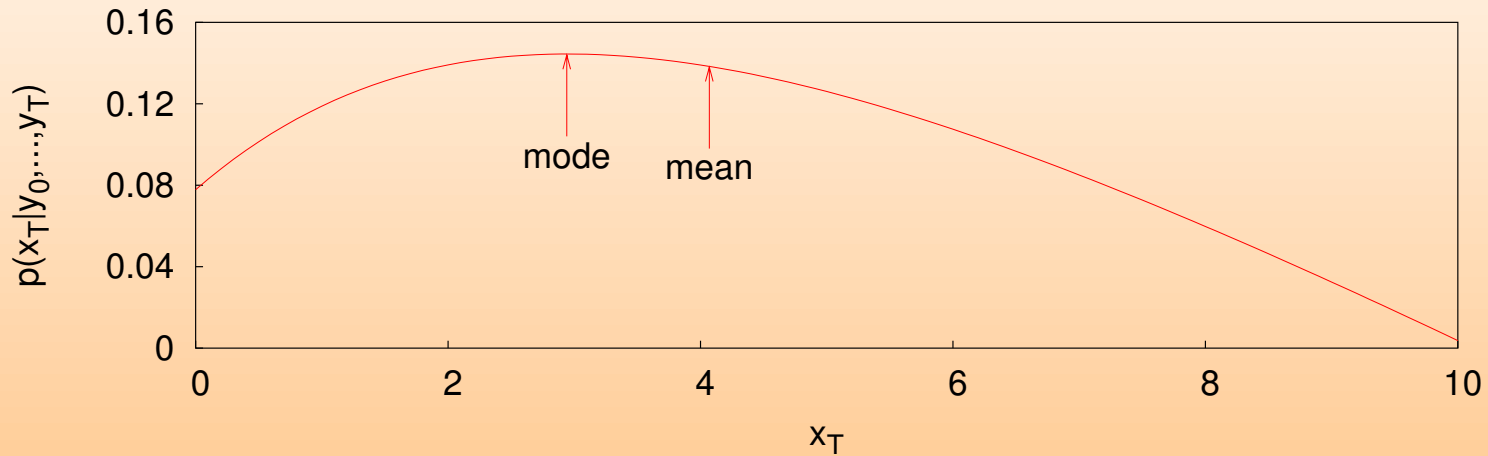
What estimate is desired?



For unconstrained linear estimation with Gaussian noise, the mean and mode of the probability distribution are the same.

Optimal estimator: Kalman filter (recursive).

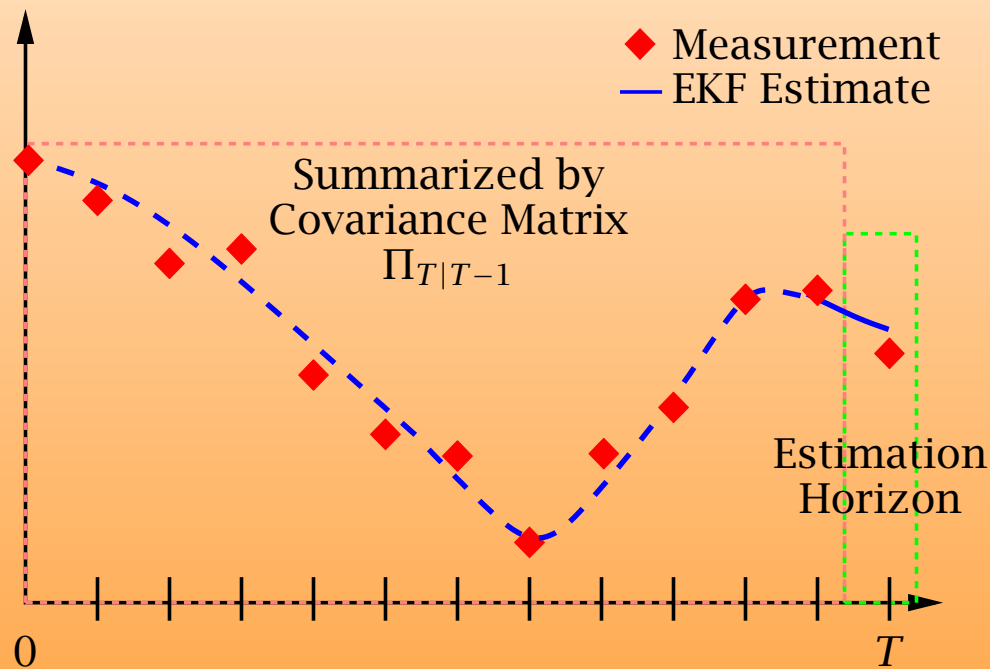
What estimate is desired?



The mean and mode of the probability distribution are generally different.

We would like to solve for the the **maximum a posteriori estimate** (MAP), i.e. the mode of this distribution.

Extended Kalman Filtering

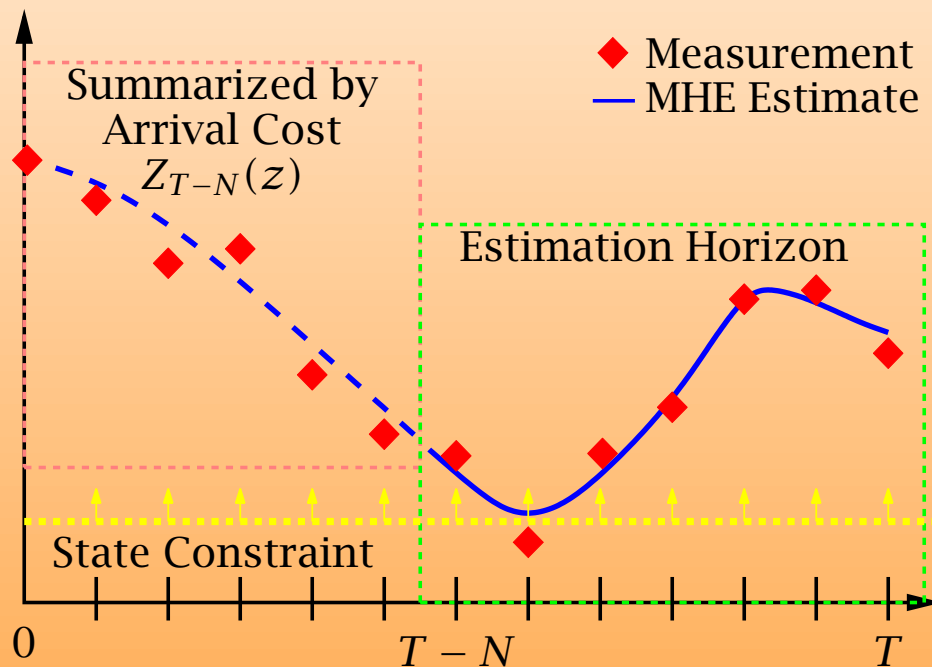


- Approximates

$$\hat{x}_{T|T} \approx \arg \max_{x_T} p(x_T | y_0, \dots, y_T)$$

- Extension of the Kalman filter to nonlinear systems via linearization
- Summarizes past data with the covariance matrix
- Computationally trivial
- Most popular industrial method

Moving Horizon Estimation



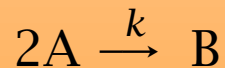
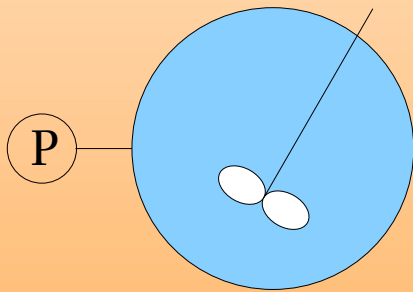
- Approximates

$$\{\hat{x}_{T-N|T}, \dots, \hat{x}_{T|T}\}$$

$$\approx \arg \max_{x_{T-N}, \dots, x_T} p(x_{T-N}, \dots, x_T | y_0, \dots, y_T)$$

- Accurately employs the non-linear model
- Can incorporate constraints
- Requires on-line optimization
- Arrival cost?

Effect of Arrival Cost: An Illustrative Example



- Well-mixed, gas phase, batch reactor
- Estimate the partial pressures of A and B
- Model

$$\frac{dx}{dt} = \begin{bmatrix} -2 & 1 \end{bmatrix}^T kP_A^2$$

- Measure the total pressure

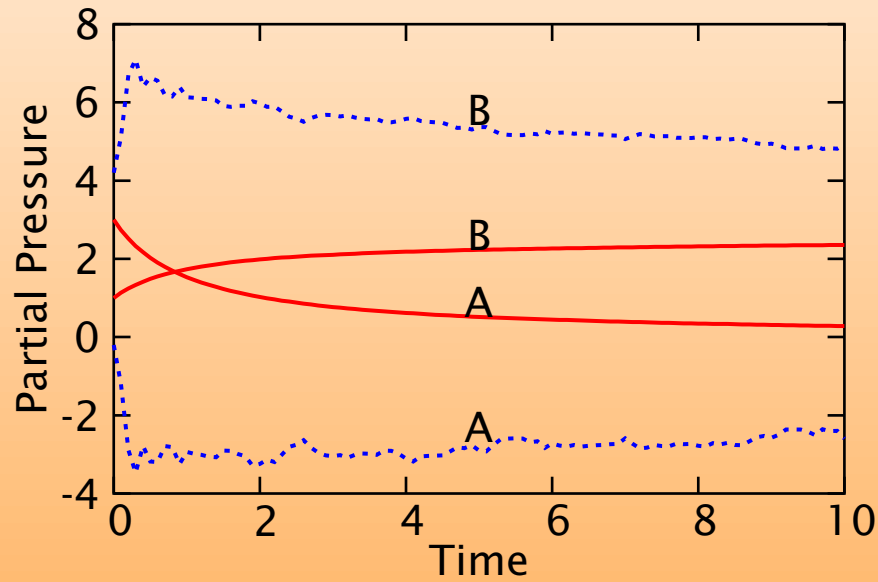
$$y = P_A + P_B$$

- Poor initial guess

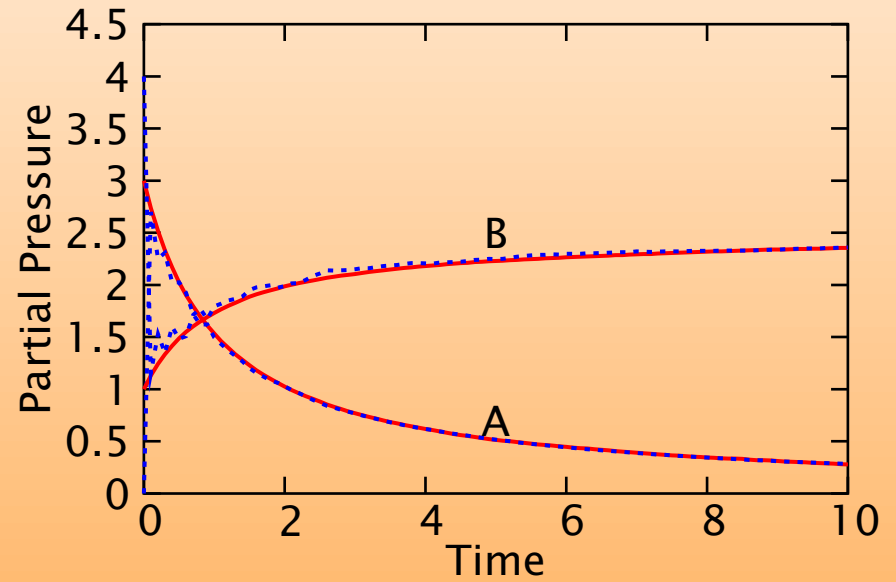
$$x_o = \begin{bmatrix} 3 & 1 \end{bmatrix}^T \text{ vs. } \bar{x}_o = \begin{bmatrix} 0.1 & 4.5 \end{bmatrix}^T$$

Estimator Comparisons

EKF



Constrained MHE



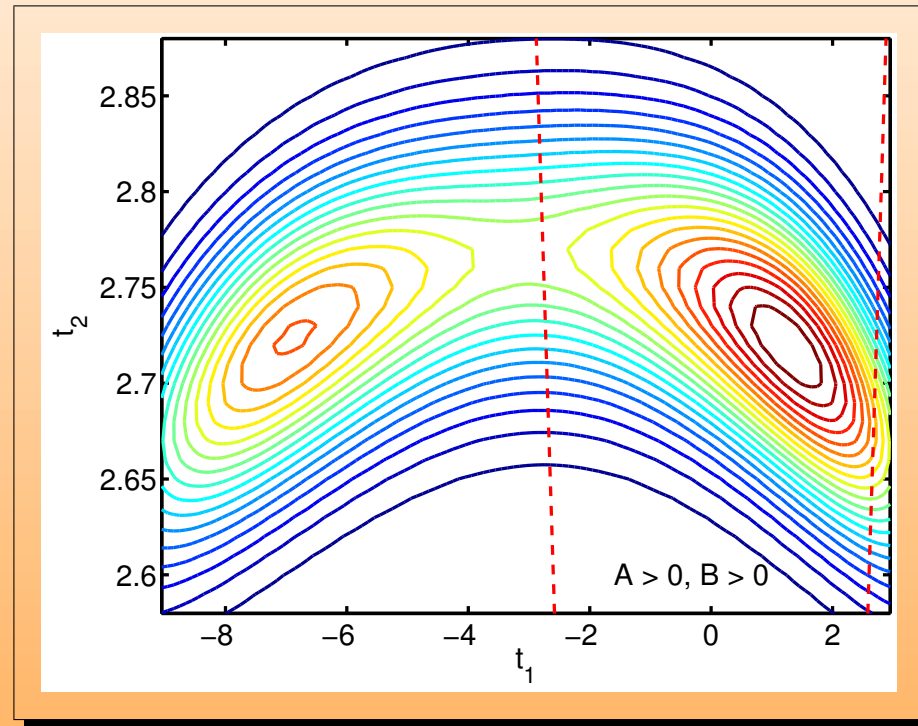
$$P_A \geq 0, P_B \geq 0$$

Horizon length of 2 minutes

Actual state (red)
Estimated state (blue)

Maximum a Posteriori Distribution

$$p(x_T | y_0, \dots, y_T)$$

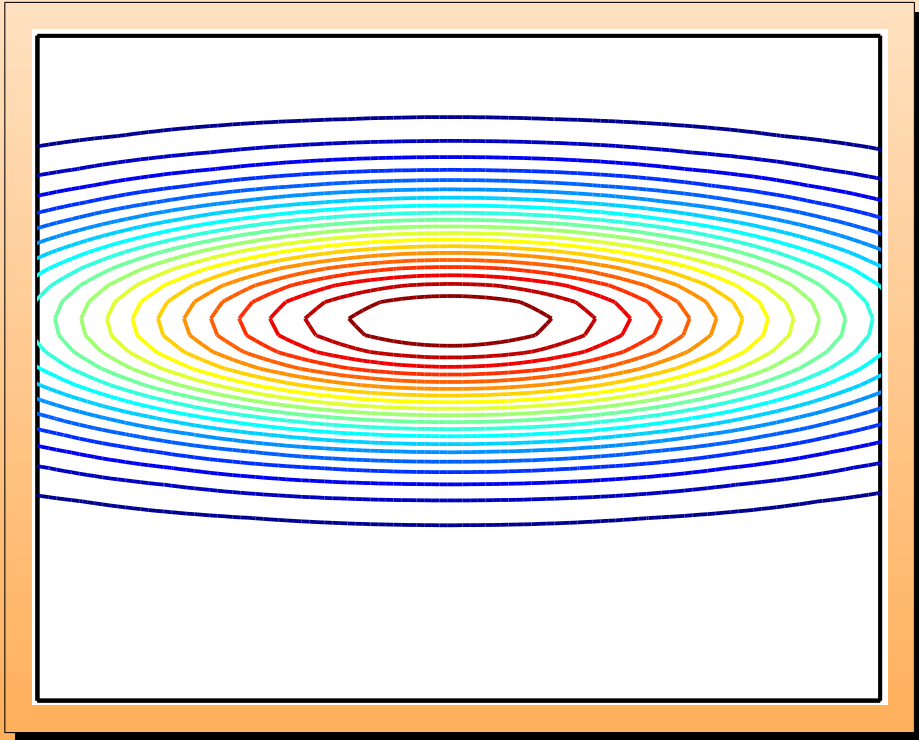


Maximum a posteriori distribution exhibits multiple optima.

Extended Kalman Filtering and Moving Horizon Estimation

Extended Kalman Filtering

$$\approx p(x_T | y_0, \dots, y_T)$$



Moving Horizon Estimation

$$\approx \max_{x_{T-N}, \dots, x_{T-1}} p(x_{T-N}, \dots, x_T | y_0, \dots, y_T)$$

Depends on the arrival cost

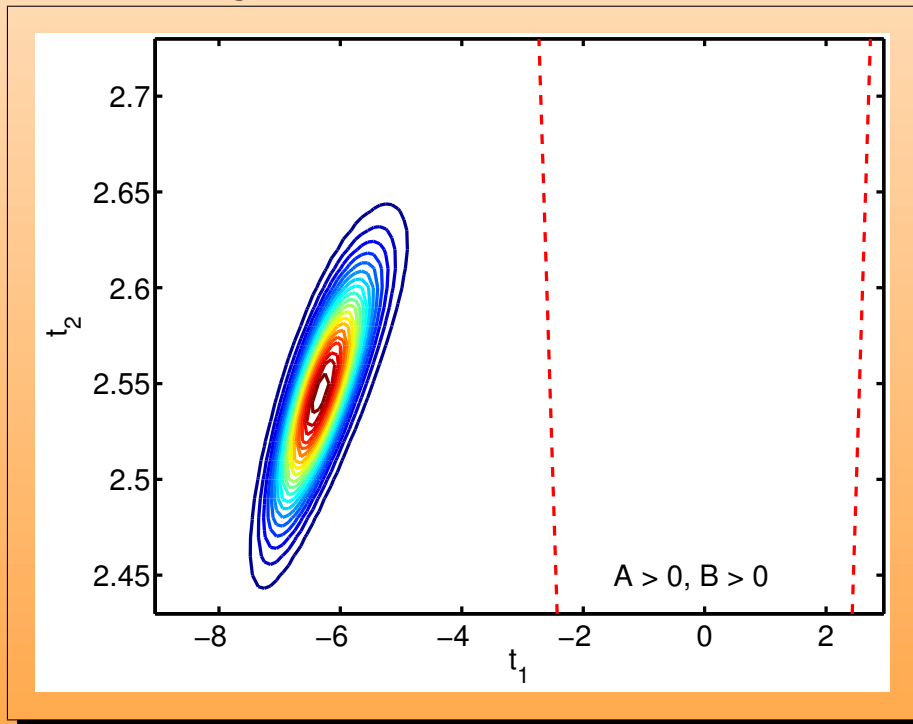
- approximate the process as a time-varying linear system?
- uniform prior?

Arrival Cost Strategies

Arrival cost approximations: smoothing update (assumes process is a time-varying linear system) or uniform prior

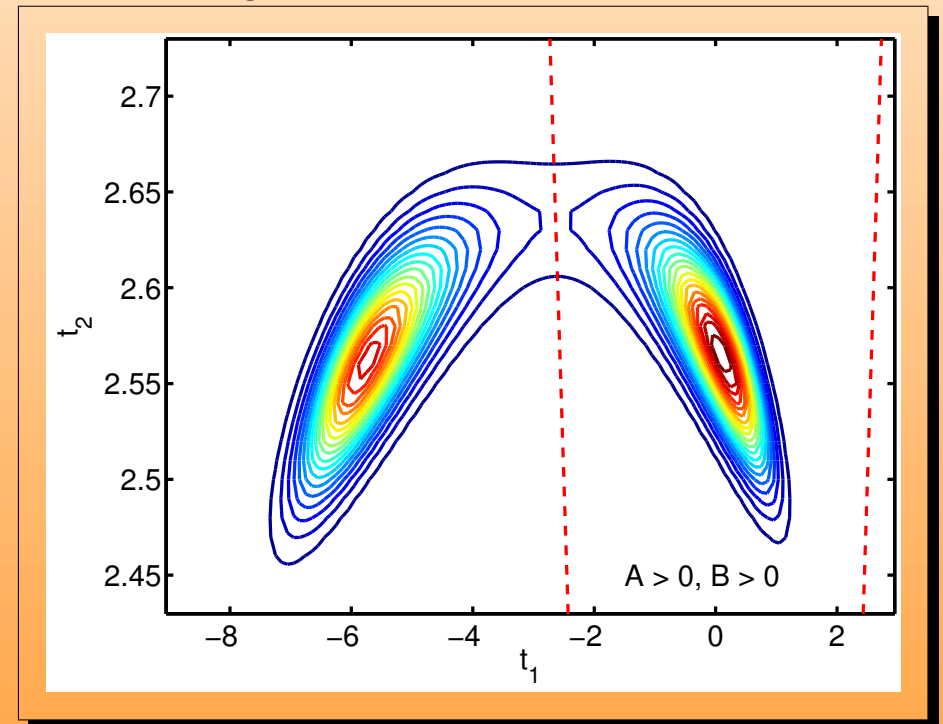
Smoothing Update

$$\max_{x_1, \dots, x_3} p(x_1, \dots, x_4 | y_0, \dots, y_4)$$



Uniform Prior

$$\max_{x_1, \dots, x_3} p(x_1, \dots, x_4 | y_0, \dots, y_4)$$



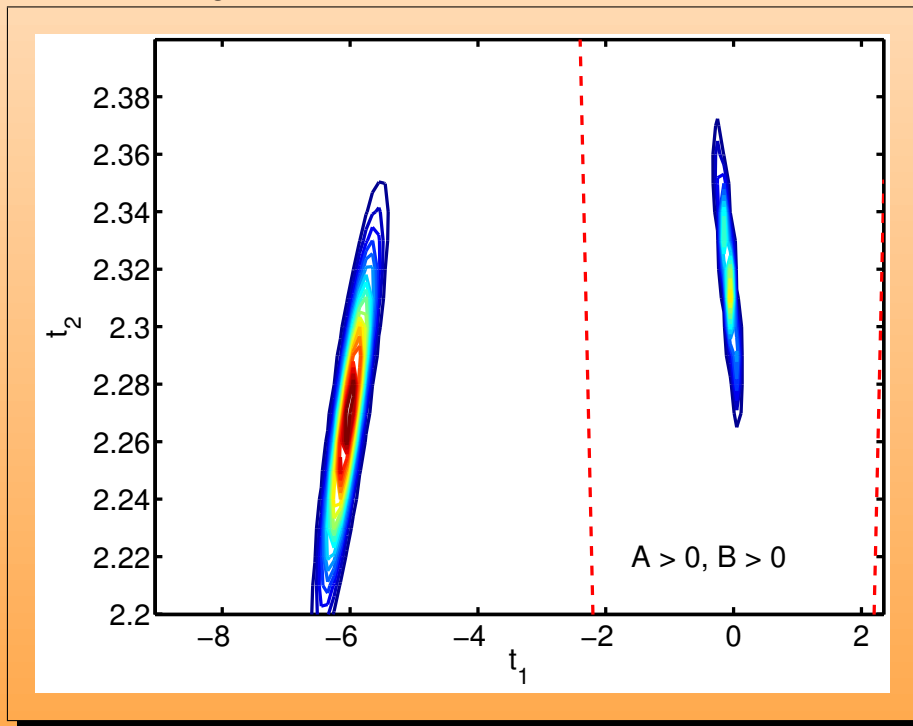
The prior (arrival cost) dominates and distorts the information contained in the data. Better to use a uniform prior if global optimization possible.

Arrival Cost Strategies: Effect of Horizon Length

Arrival cost approximations: smoothing update (assumes process is a time-varying linear system) or uniform prior

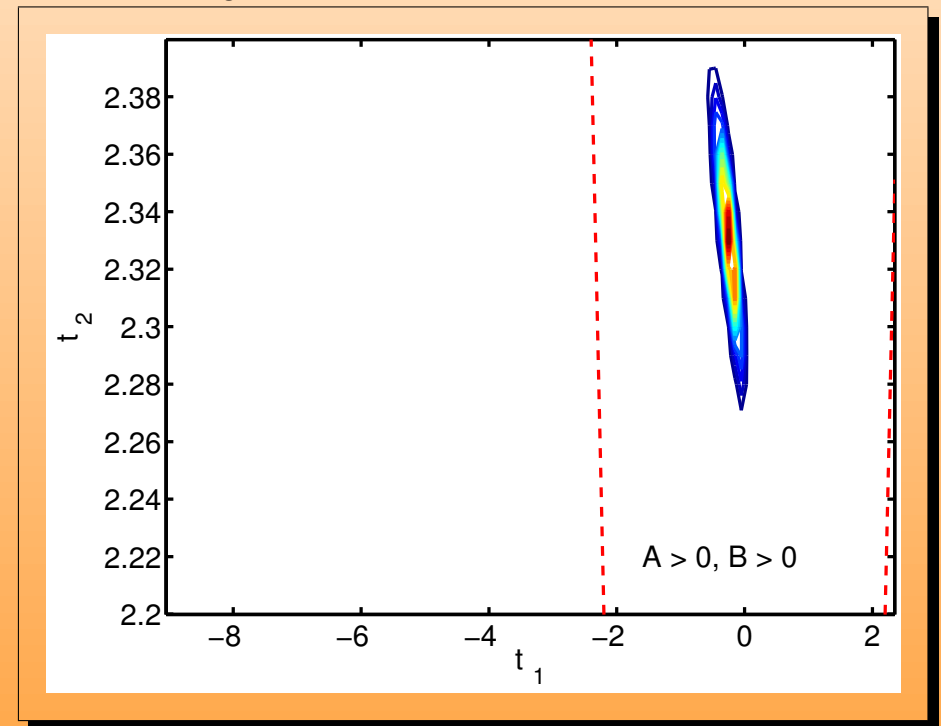
Smoothing Update

$$\max_{x_1, \dots, x_9} p(x_1, \dots, x_{10} | y_0, \dots, y_{10})$$



Uniform Prior

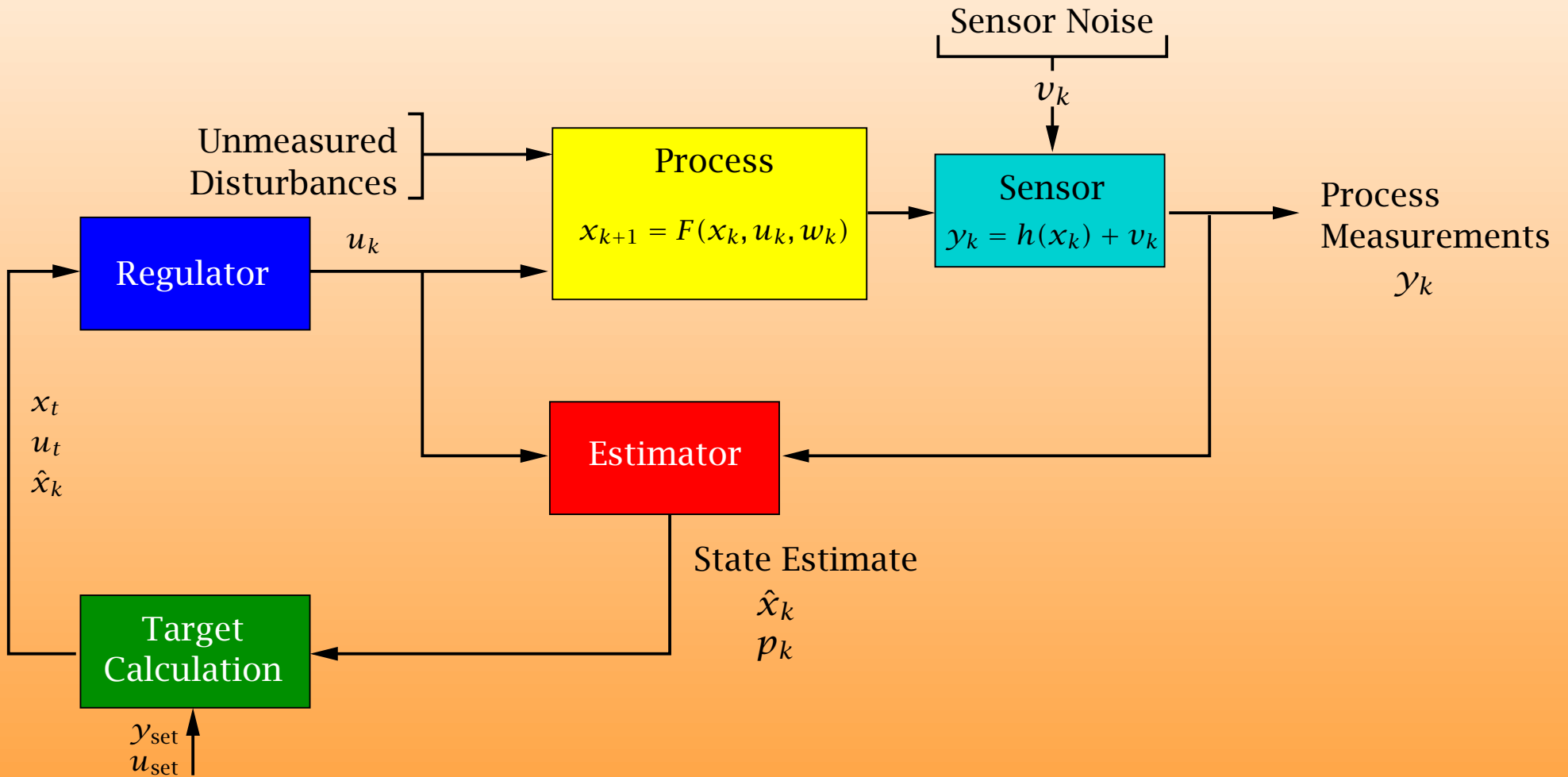
$$\max_{x_1, \dots, x_9} p(x_1, \dots, x_{10} | y_0, \dots, y_{10})$$



Arrival Cost Conclusions

- The EKF can predict only one optimum.
Estimation behavior dictated by initial region of attraction.
- Behavior of MHE depends upon the arrival cost.
 - Approximating the past behavior as a time-varying, linear system introduces significant bias
 - Longer horizon can overcome effects of poor arrival cost
 - Best (current) option: uniform prior with global optimization

Closed-Loop Control



We will use the NMPC toolbox for the estimator, regulator, and target calculation (local optimization) [3].

Disturbance models for nonlinear models

- For offset free control, must account for discrepancies between the plant and the model.
- Augment the state with a disturbance model:

$$\mathbf{x}_{k+1} = F(\mathbf{x}_k, \mathbf{u}_k + \mathbf{X}_u \mathbf{d}_k) + G \mathbf{w}_k$$

$$y_k = h(\mathbf{x}_k) + \mathbf{X}_y \mathbf{d}_k + v_k$$

$$\mathbf{d}_{k+1} = \mathbf{d}_k + \boldsymbol{\xi}_k$$

$$\boldsymbol{\xi}_k \sim \mathcal{N}(0, Q_d)$$

Implies that \mathbf{d}_k is stochastic!

Plant-model mismatch: exothermic CSTR example

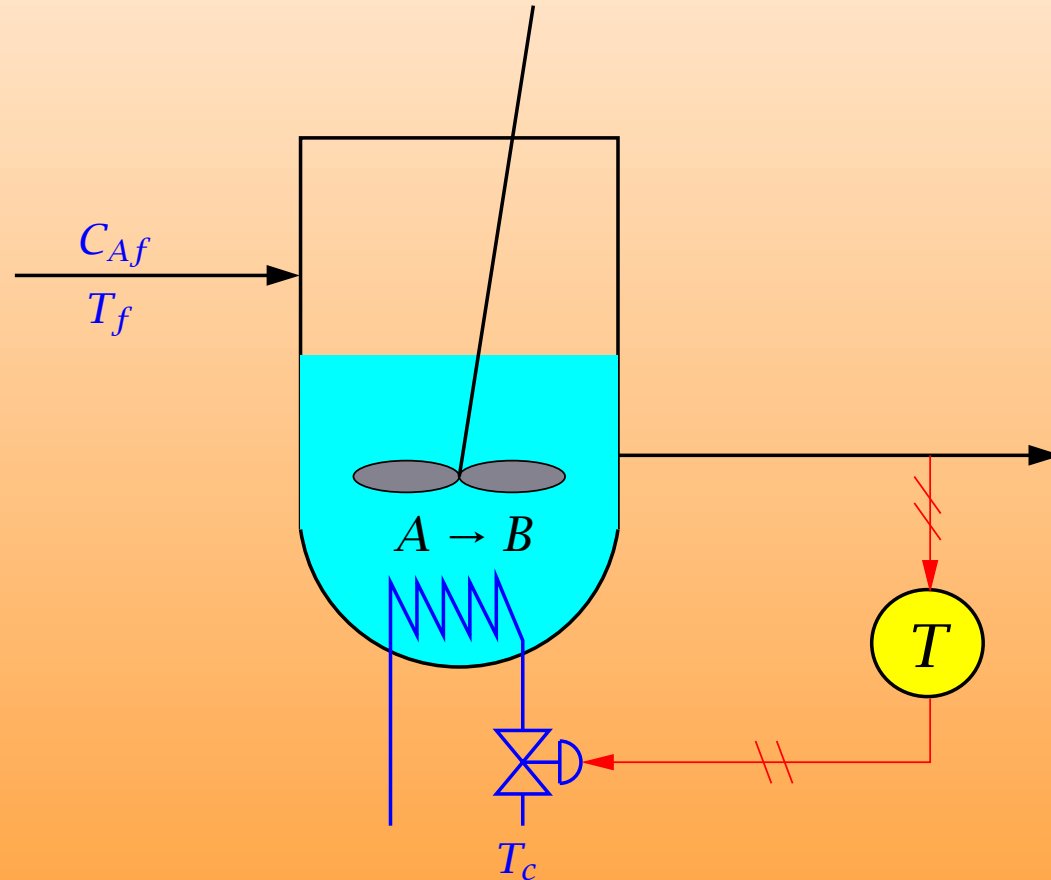
$$x = \begin{bmatrix} c_A \\ T \end{bmatrix}$$

$$u = T_c$$

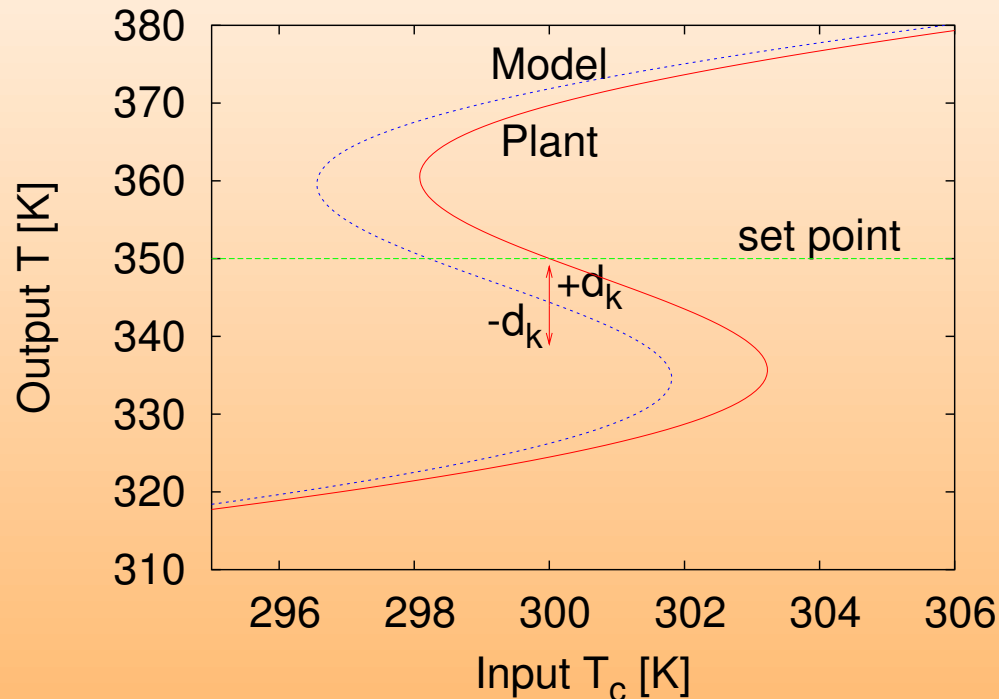
$$|\Delta u_k| \leq 15 \text{ K}$$

$$y_k = \begin{bmatrix} 0 & 1 \end{bmatrix} x_k + d_k$$

Small mismatch in activation energy between the plant and the model.



Output disturbance model generates multiple steady states!



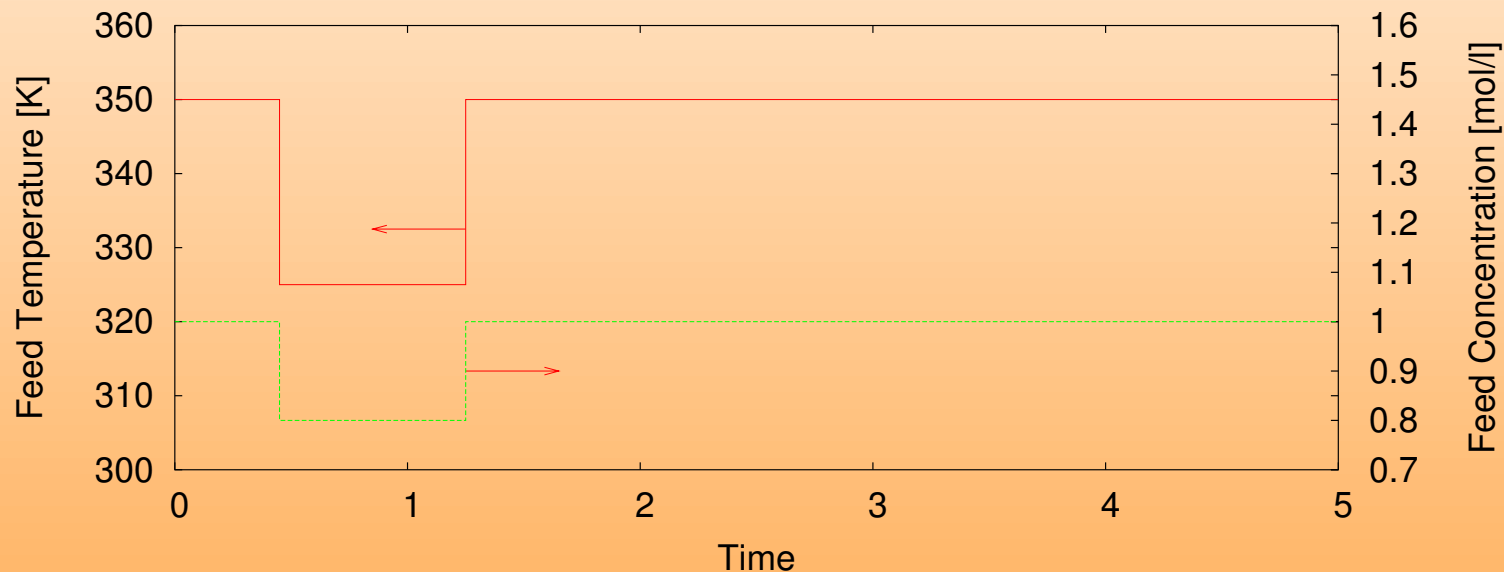
Multiple optima arise in the estimator. Do these optima affect control performance?

	c_A (mol/l)	Output T (K)	Disturbance d (K)
1	0.851	326.2	23.8
2	0.583	344.4	5.6
3	0.177	371.8	-21.8

Model Steady States for a Plant with $T_c = 300$ K, $T = 350$ K

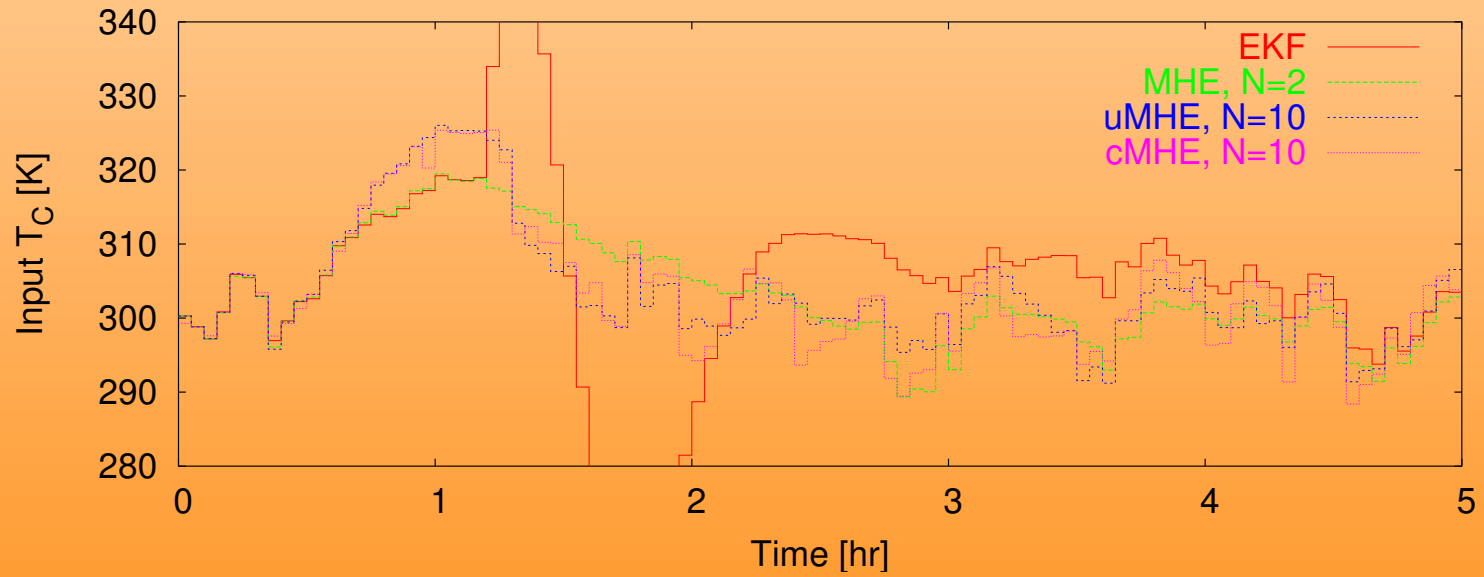
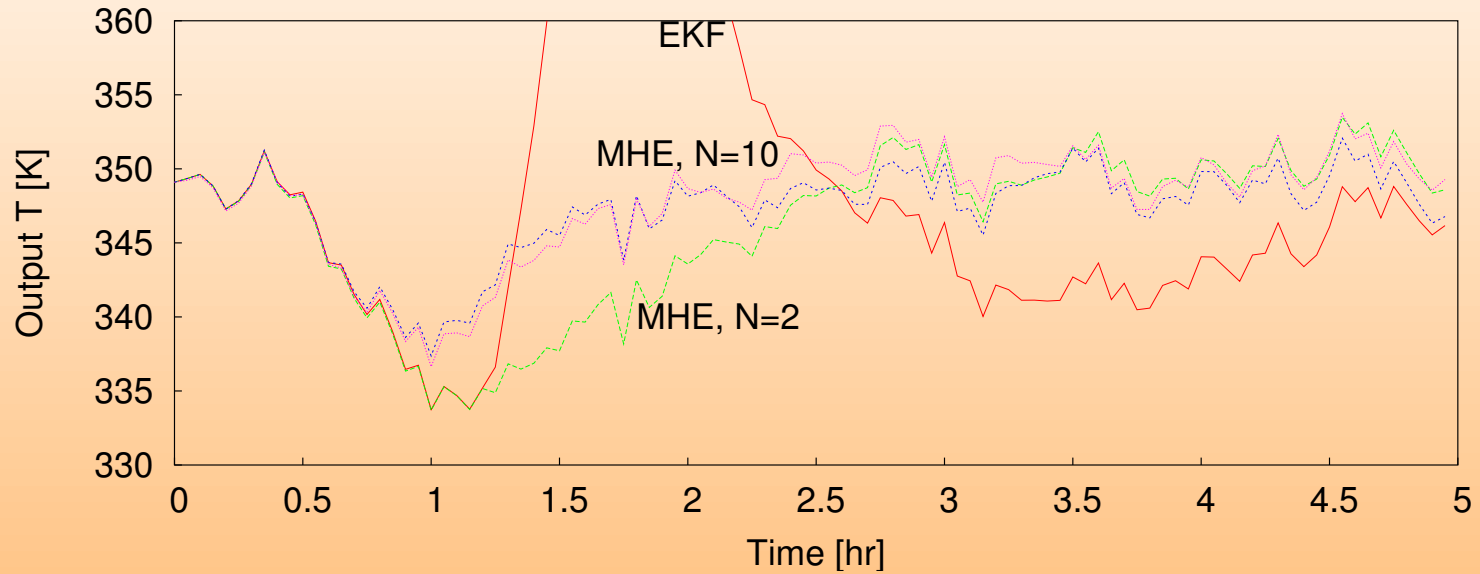
Example: Disturbance rejection

- Consider a disturbance in the feed:

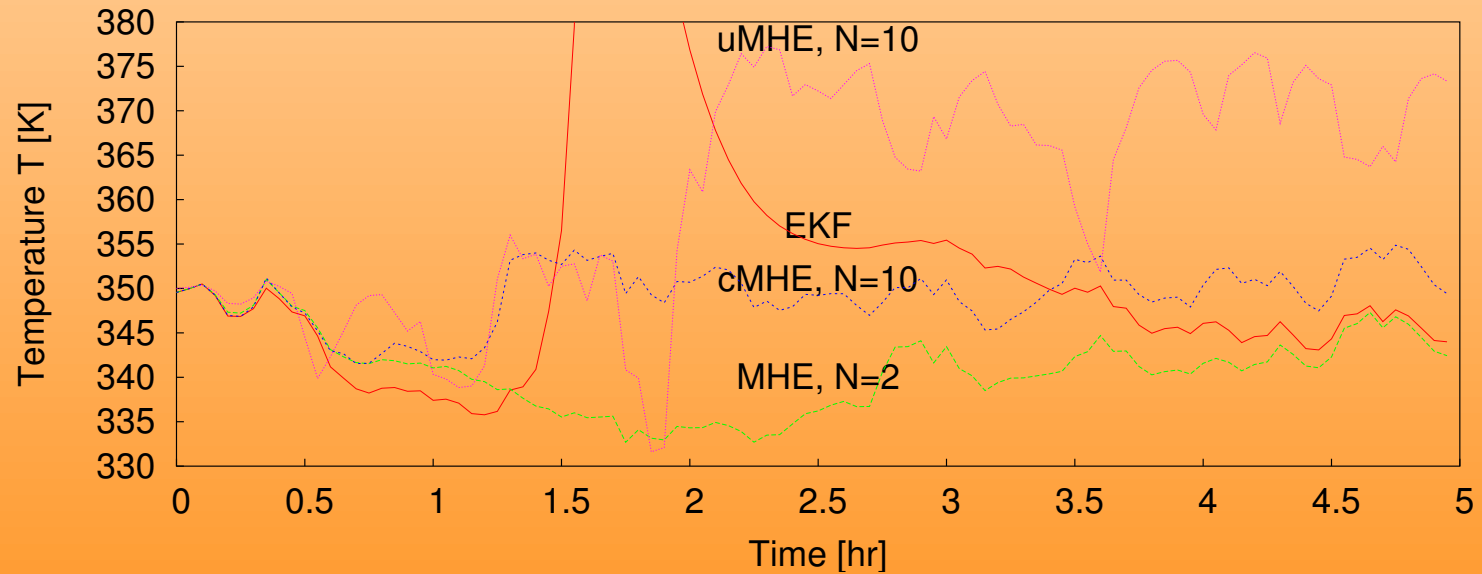
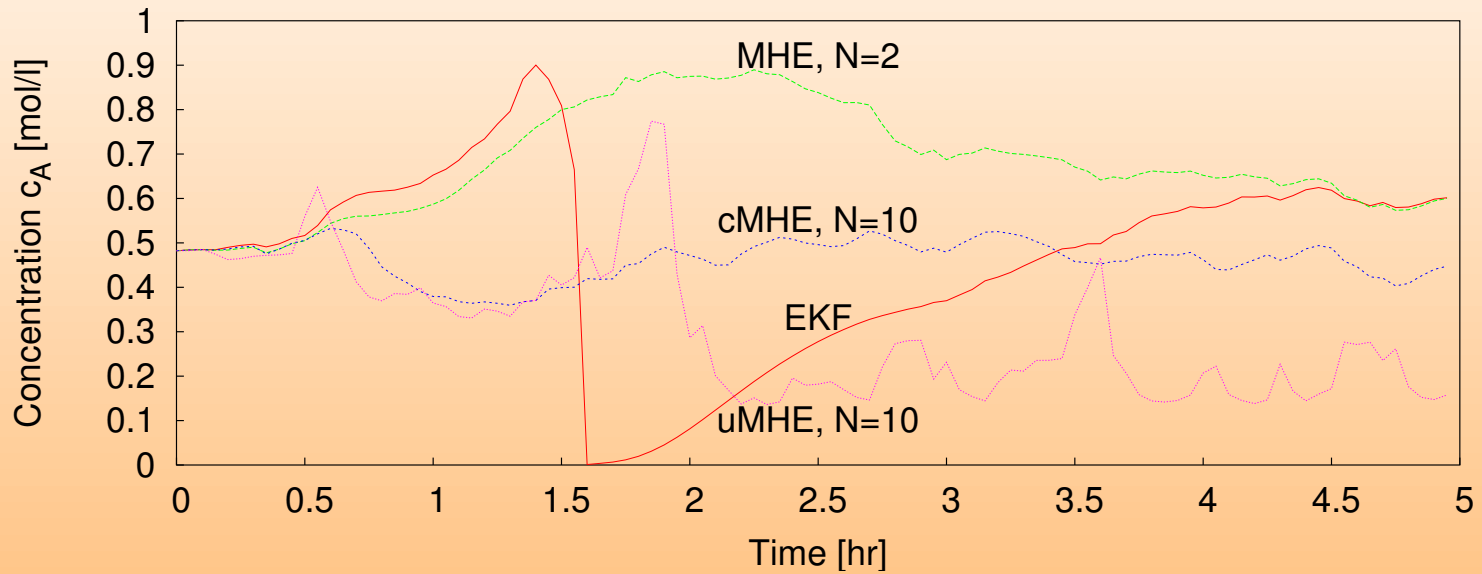


- Estimators:
 1. EKF
 2. MHE with $N = 2$, smoothing update
 3. MHE with $N = 10$, no initial penalty
 4. MHE with $N = 10$, constant initial penalty

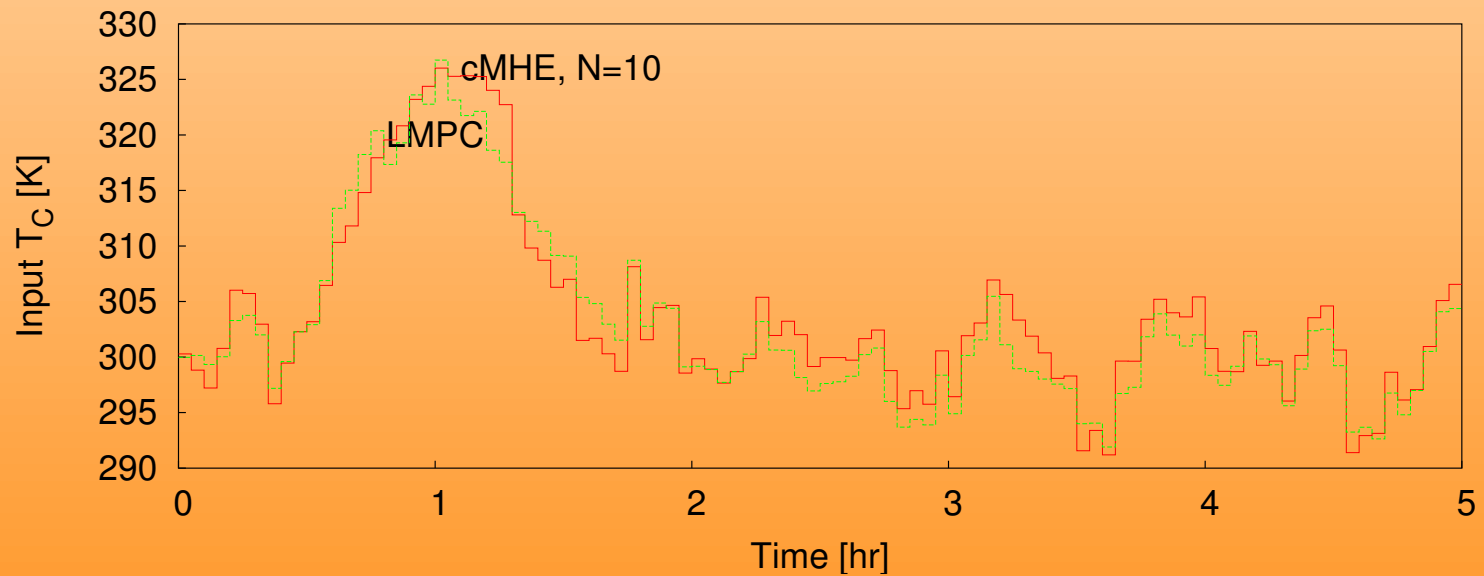
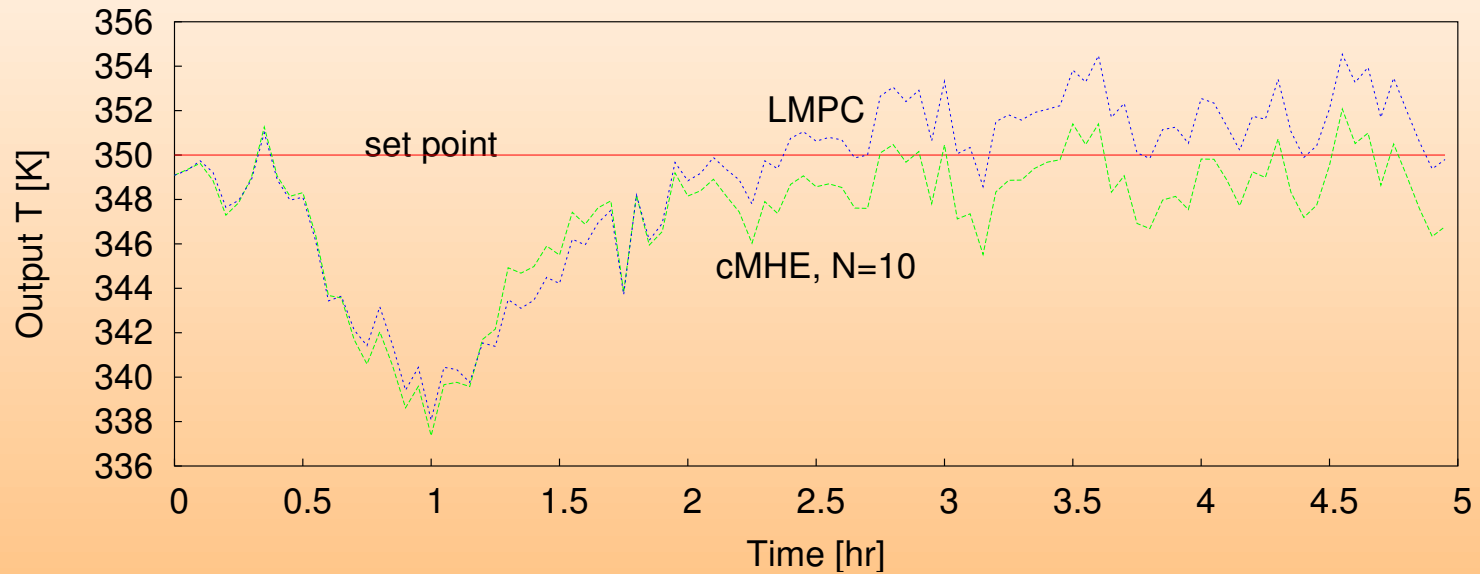
Exothermic CSTR Results



Exothermic CSTR Results



Comparison to Linear MPC Results



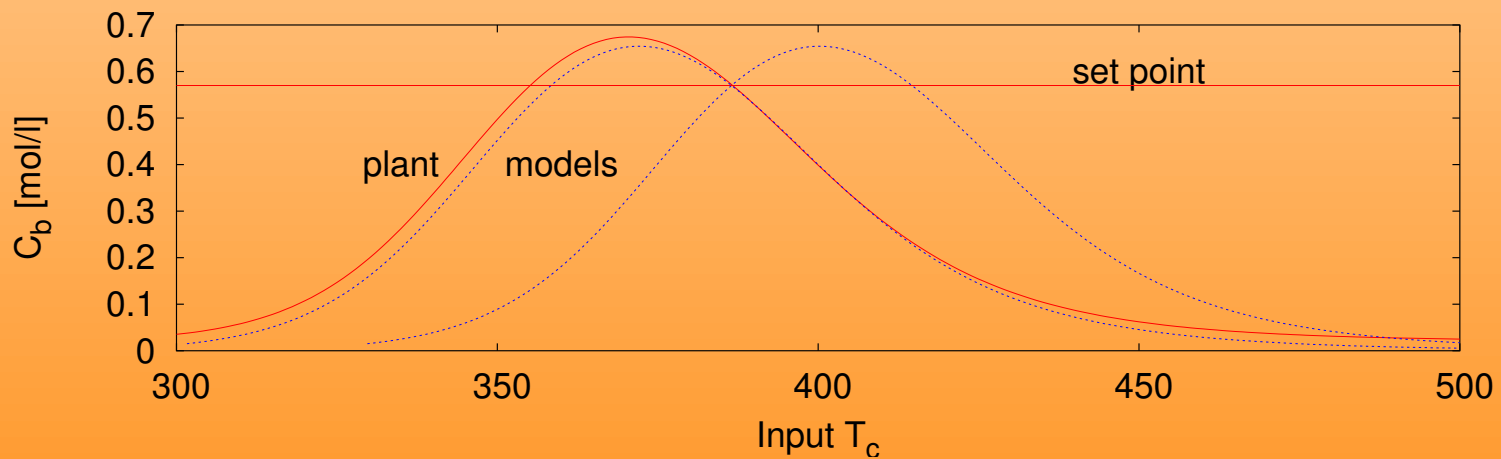
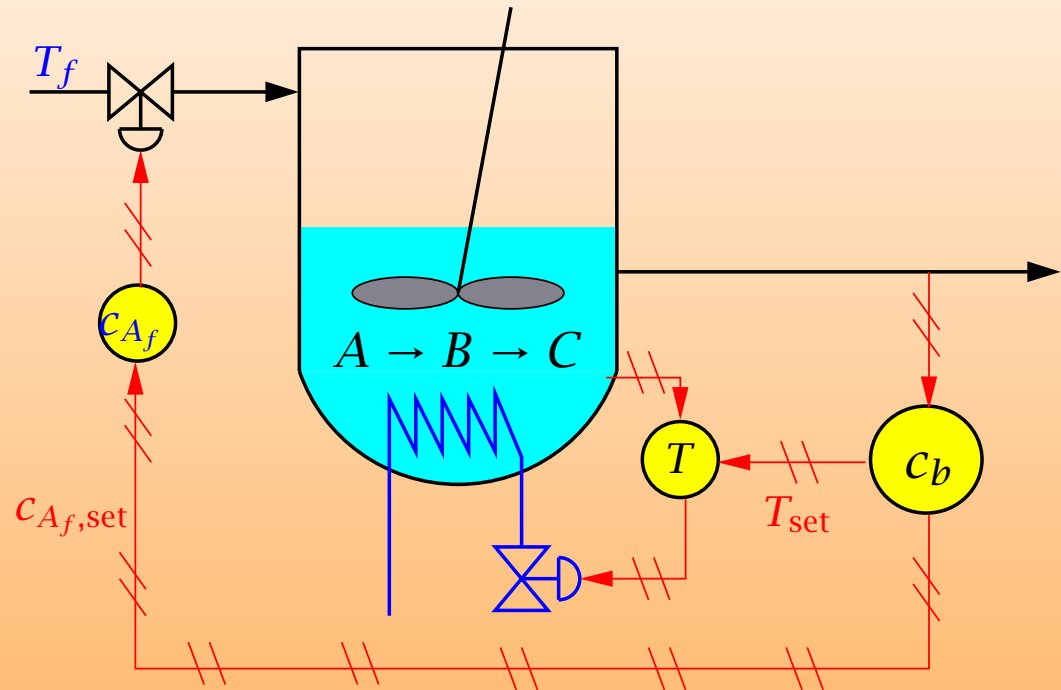
Plant-model mismatch: maximum yield example

$$\mathbf{x} = \begin{bmatrix} c_A & c_B \end{bmatrix}^T$$

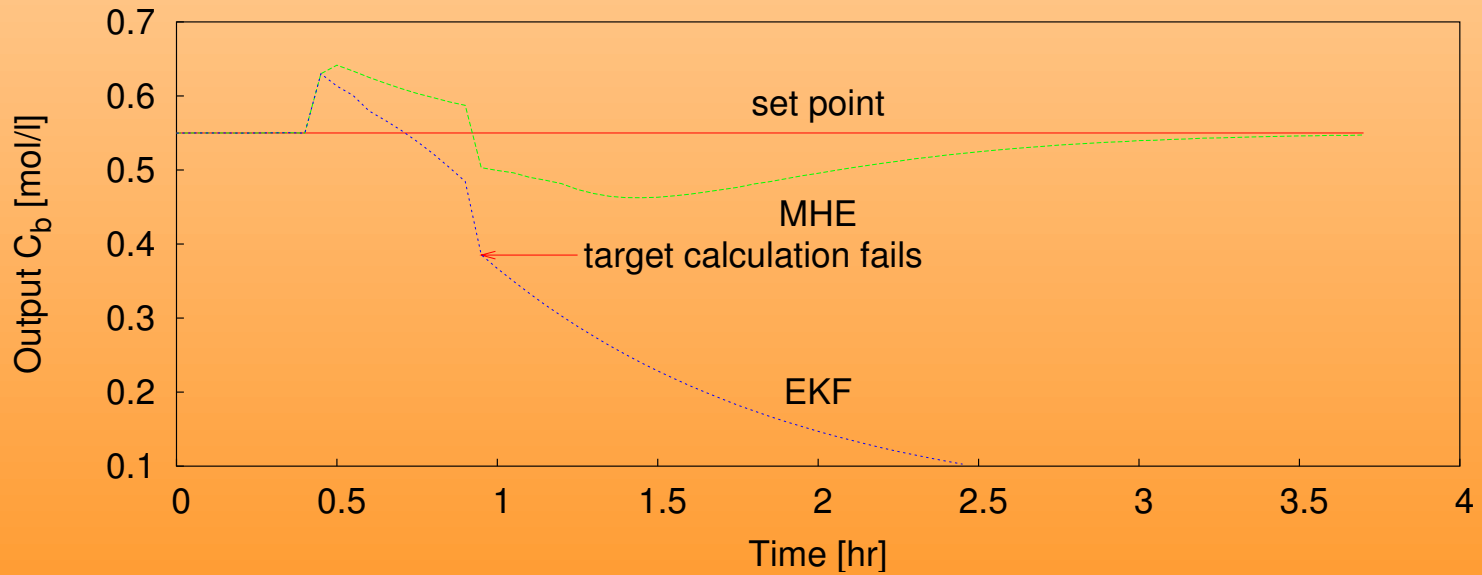
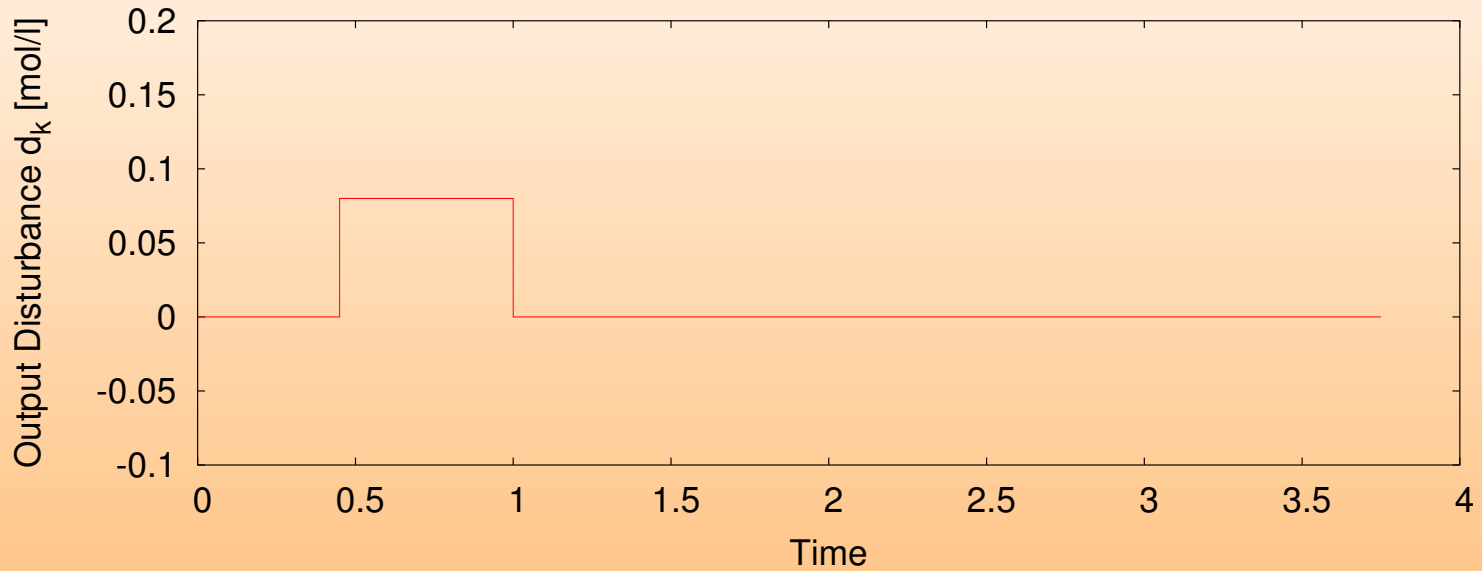
$$\mathbf{u} = \begin{bmatrix} T_c & c_{Af} \end{bmatrix}^T$$

$$\mathbf{u}_k = \begin{bmatrix} T_c \\ c_{Af} \end{bmatrix} + \begin{bmatrix} 1 \\ 0 \end{bmatrix} d_k$$

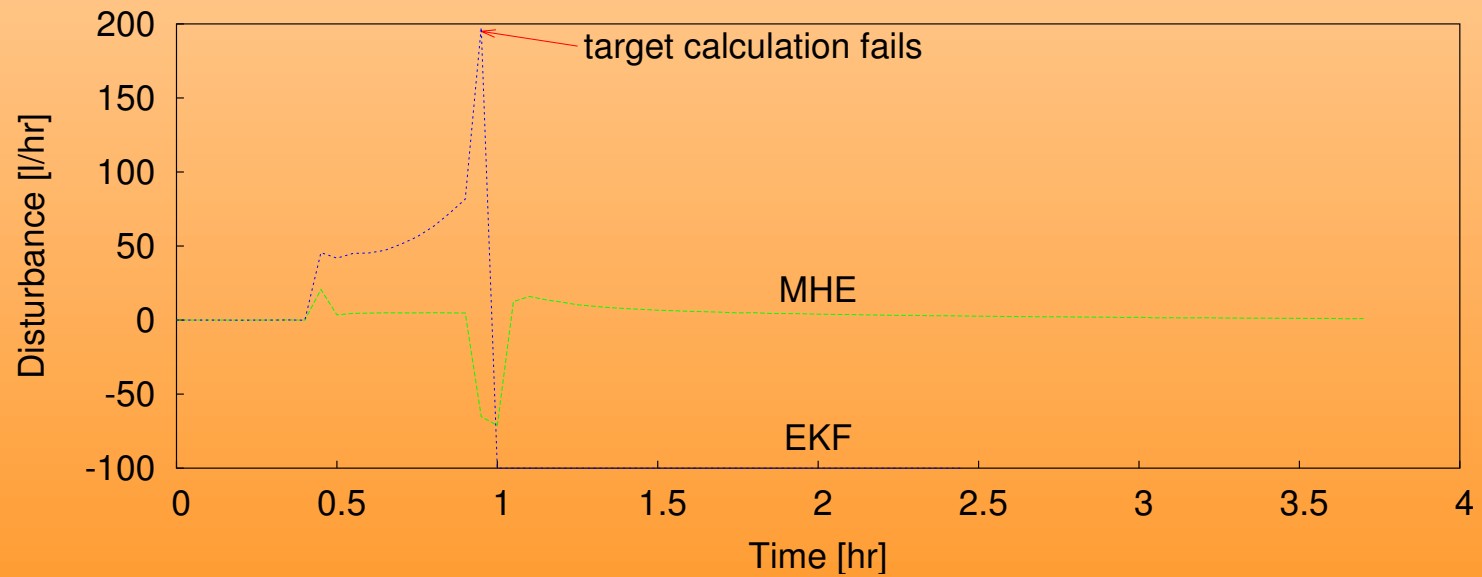
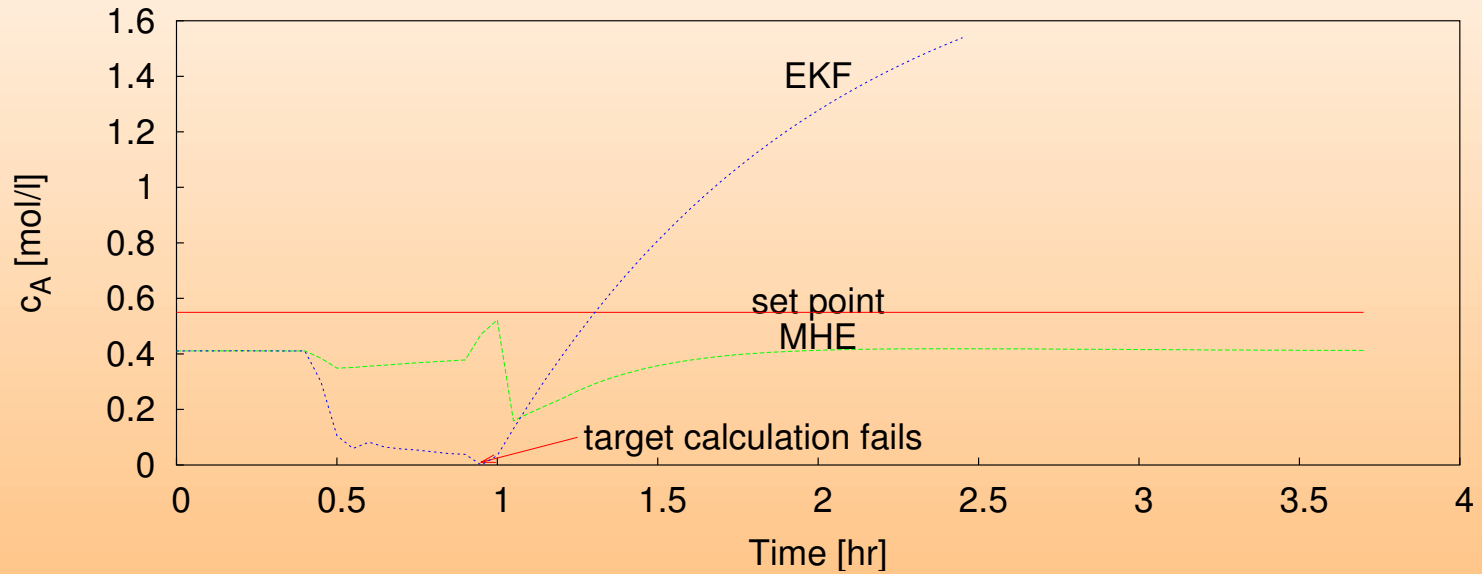
$$\mathbf{y}_k = \begin{bmatrix} 0 & 1 \end{bmatrix} \mathbf{x}_k$$



Maximum Yield CSTR Results



Maximum Yield CSTR Results



Conclusions

- Integrated disturbance models can induce multiple optima with nonlinear models.
- EKF: poor state tracking results in poor control.
- MHE
 1. better control than the EKF
 2. increased horizon length results in improved performance
- No significant improvement in nonlinear over linear control for disturbance rejection.
- If you want better disturbance rejection, you need a better disturbance model.

Acknowledgements

- Organizers of the Gordon Research Conference for Statistics in Chemistry & Chemical Engineering
- Wen-shiang Chen and Prof. Bhavik Bakshi (Ohio State University)

Questions?

Improving MHE

- Primary concern: longer horizon length = greater computational expense
- Basis of MHE:

$$\begin{aligned} & \max_{x_{T-N+1}, \dots, x_T} p(x_{T-N+1}, \dots, x_T | y_0, \dots, y_T) \\ &= \max_{x_{T-N+1}, \dots, x_T} \underbrace{p(x_{T-N+1} | y_0, \dots, y_T)}_{\text{Arrival Cost}} \underbrace{\left(\prod_{k=T-N+1}^{T-1} p(x_{k+1} | x_k) \right) \left(\prod_{k=T-N+1}^T p(y_k | x_k) \right)}_{\text{Estimation Horizon}} \end{aligned}$$

- If we had a better estimate for $p(x_{T-N+1} | y_0, \dots, y_T)$, we could shorten the estimation horizon
- Can we use Monte Carlo filters to estimate this density?

State Estimation via Monte Carlo Filters

- Basic idea: reconstruct state estimates from simulations of the stochastic process

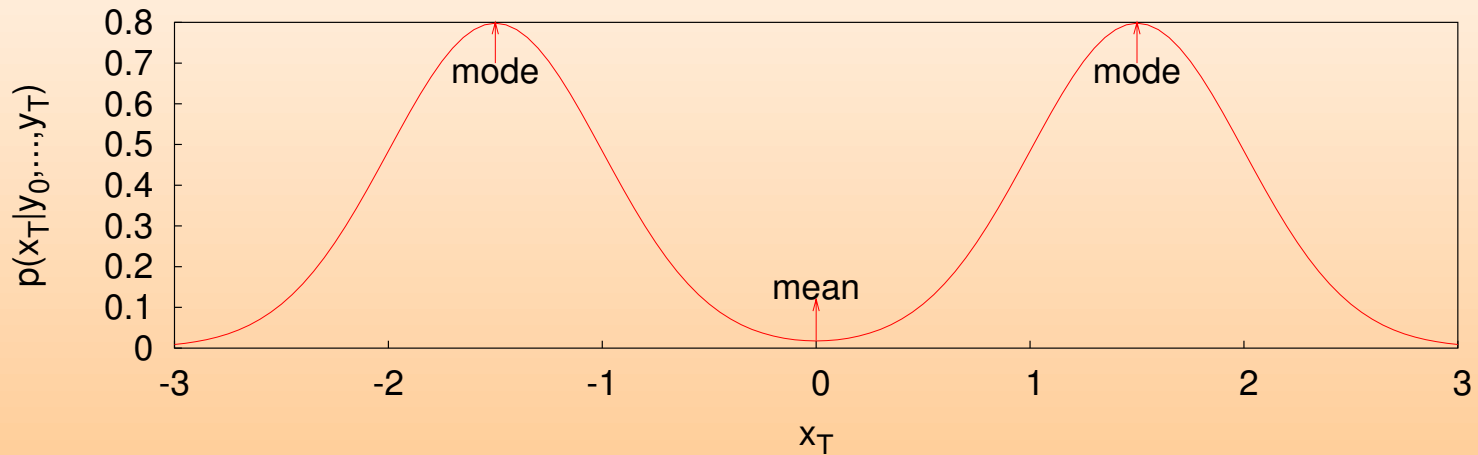
$$\int h(\mathbf{x})P(\mathbf{x})d\mathbf{x} = \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N h(\mathbf{x}^i)$$

- Most MC filters propose estimation of the mean

$$E[\mathbf{x}] = \int \mathbf{x}P(\mathbf{x})d\mathbf{x} \approx \frac{1}{N} \sum_{i=1}^N \mathbf{x}^i$$

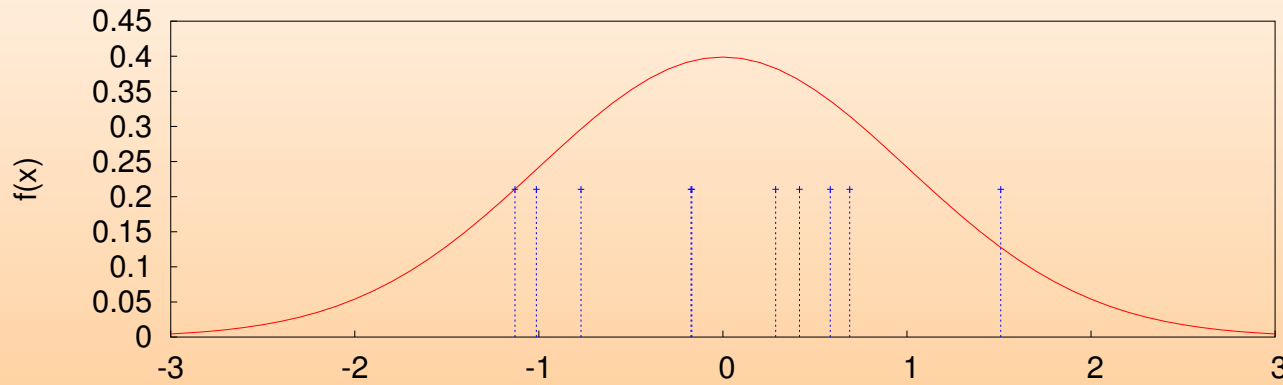
- Permits use of any combination of model, random noise
- We will consider rejection sampling [1]

What estimate is desired?

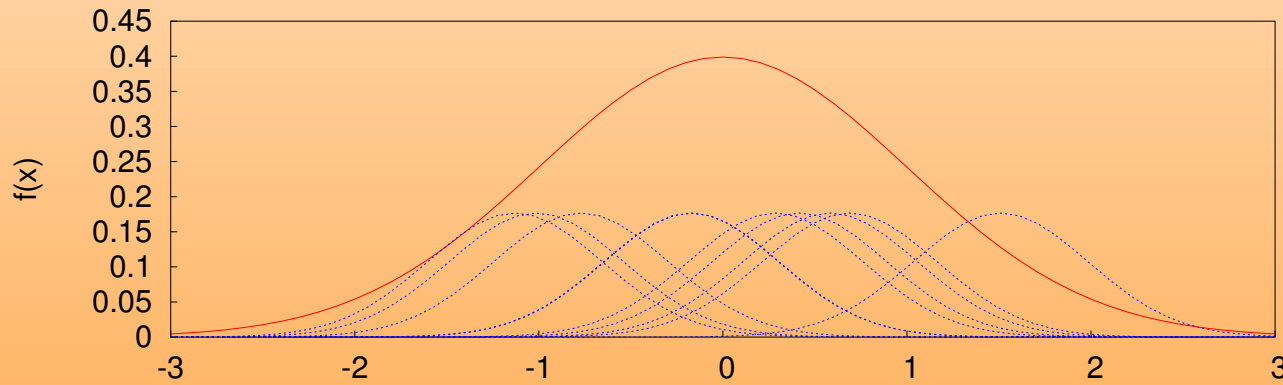


- MHE estimates the mode
- Monte Carlo filters must estimate the entire probability density to calculate the mode
 1. density estimation
 2. optimization

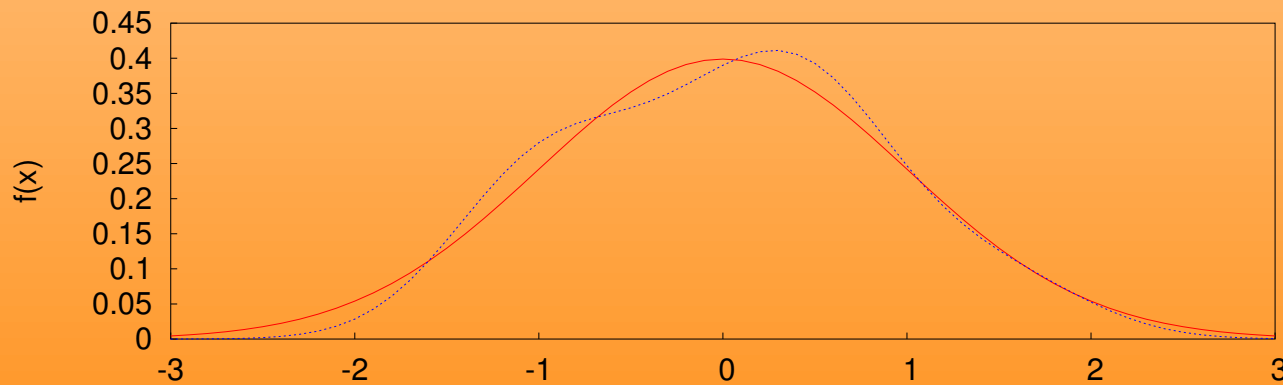
Example: density estimation of a normal distribution



Draw samples from the underlying distribution.

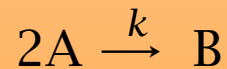
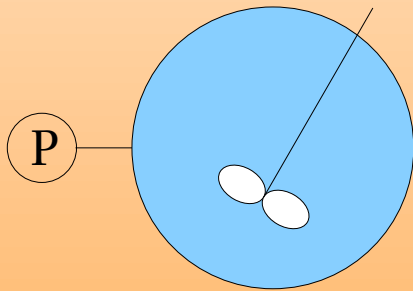


Apply a symmetric “kernel” density at each sample.



Sum the kernel densities to approximate the underlying distribution.

A Simple Example



- Well-mixed, gas phase, batch reactor
- Estimate the partial pressures of A and B
- Model

$$\frac{dx}{dt} = \begin{bmatrix} -2 & 1 \end{bmatrix}^T kP_A^2$$

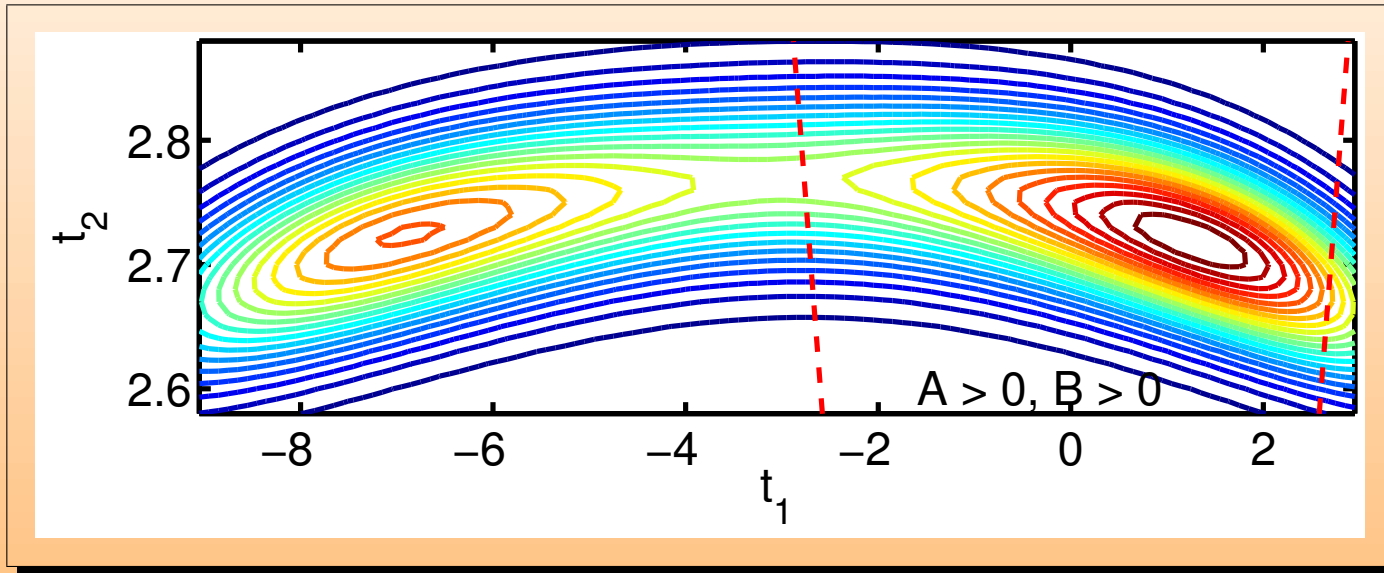
- Measure the total pressure

$$y = P_A + P_B$$

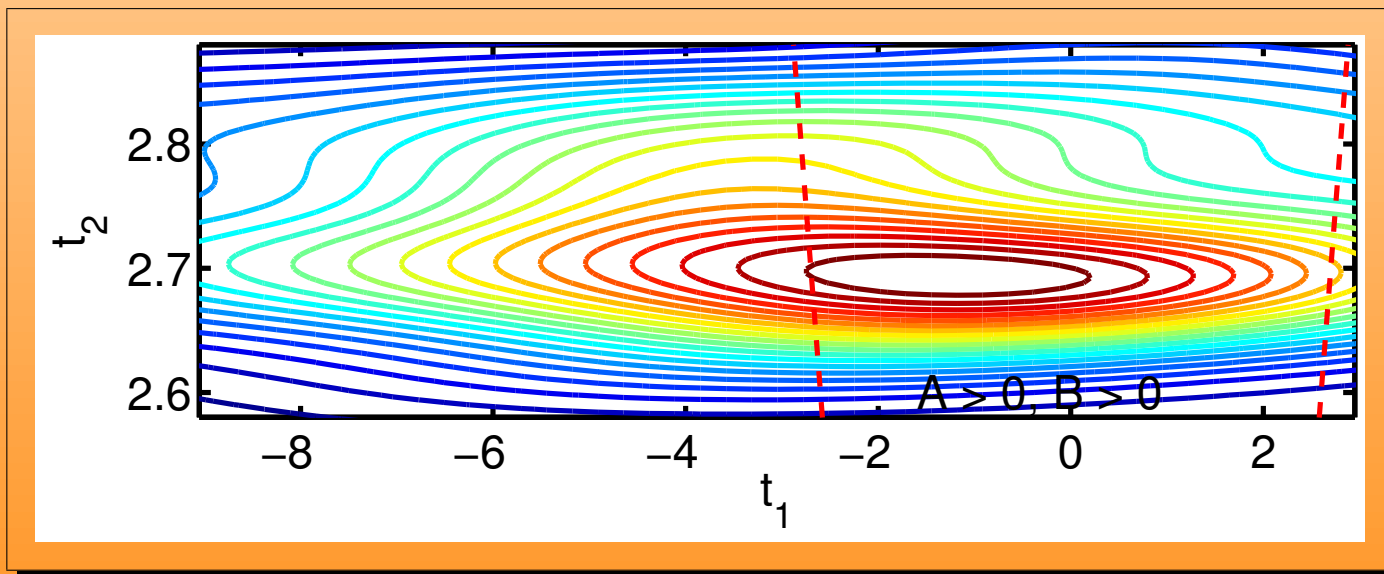
- Poor initial guess

$$x_o = \begin{bmatrix} 3 & 1 \end{bmatrix}^T \text{ vs. } \bar{x}_o = \begin{bmatrix} 0.1 & 4.5 \end{bmatrix}^T$$

A Posteriori Comparison: Actual vs. Monte Carlo $p(x_1|y_0, y_1)$



Actual



Monte Carlo
Reconstruction
(100 Accepted
Samples)

Comments on Monte Carlo Estimation

- Not very accurate estimation of the mode.

Sources of error:

1. finite number of samples
 2. density estimation approximation
- Simple to code
 - May provide a useful estimate of the arrival cost if computationally inexpensive

Density estimation: the curse of dimensionality [2]

Dimensionality	Required Sample Size
1	4
2	19
3	67
4	223
5	768
6	2790
7	10700
8	43700
9	187000
10	842000

Sample size required to ensure that the relative mean square error at zero (a single point) is less than 0.1. The underlying distribution is a standard multivariate normal density.

References

- [1] E. Bølviken, P. J. Acklam, N. Christopherson, and J.-M. Størdal. Monte Carlo filters for non-linear state estimation. *Automatica*, 37(2):177–183, February 2001.
- [2] B. W. Silverman. *Density Estimation for Statistics and Data Analysis*. Chapman and Hall, New York, 1986.
- [3] M. Tenny. *Computational Strategies for Nonlinear Model Predictive Control*. PhD thesis, University of Wisconsin–Madison, 2002.